

UNIVERSITI SAINS MALAYSIA

First Semester Examination
Academic Session 2004/2005

October 2004

**MST 561 – STATISTICAL INFERENCE
[PENTAABIRAN STATISTIK]**

Duration : 3 hours

[Masa : 3 jam]

Please check that this examination paper consists of **NINE [9]** pages of printed material before you begin the examination.

*[Sila pastikan bahawa kertas peperiksaan ini mengandungi **SEMBILAN [9]** muka surat yang bercetak sebelum anda memulakan peperiksaan ini.]*

Answer all **FIVE** questions.

Jawab semua **LIMA** soalan.

1. (a) Let X_1, X_2, \dots, X_n be independent, identically distributed random variables with common probability density function.

$$f(x) = \begin{cases} \theta e^{-\theta x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

Let $Y_1 = X_{(1)}, Y_2 = X_{(2)}, \dots, Y_n = X_{(n)}$ be the order statistics for X_1, X_2, \dots, X_n .

Let $Z_1 = nY_1, Z_2 = (n-1)(Y_2 - Y_1), \dots, Z_n = (Y_n - Y_{n-1})$.

Show that Z_1, Z_2, \dots, Z_n are independent and identically distributed random variables.

[50 marks]

- (b) Let X_1, X_2, \dots, X_n be a random sample of size n from a Poisson distribution with parameter $\lambda = 1$, i.e.

$$P(X_i = x) = \frac{e^{-1}}{x!}, \quad x = 0, 1, 2, \dots, i = 1, 2, \dots, n$$

(i) Derive the m.g.f. for X_i .

(ii) Let $\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}$. Consider

$Y_n = \sqrt{n}(\bar{X}_n - 1)$. Show that the m.g.f for Y_n is given by $\exp \left[-t\sqrt{n} + n(e^{t/\sqrt{n}} - 1) \right]$

(iii) Hence, obtain the limiting distribution of Y_n when $n \rightarrow \infty$.

[50 marks]

1. (a) Katakan X_1, X_2, \dots, X_n adalah pembolehubah-pembolehubah rawak takbersandar dan tertabur secaman dengan fungsi ketumpatan kebarangkalian sepunya

$$f(x) = \begin{cases} \theta e^{-\theta x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

Katakan $Y_1 = X_{(1)}, Y_2 = X_{(2)}, \dots, Y_n = X_{(n)}$ adalah statistik tertib bagi X_1, X_2, \dots, X_n .

Katakan $Z_1 = nY_1, Z_2 = (n-1)(Y_2 - Y_1), \dots, Z_n = (Y_n - Y_{n-1})$.

Tunjukkan bahawa Z_1, Z_2, \dots, Z_n adalah pembolehubah-pembolehubah rawak takbersandar dan tertabur secaman.

[50 markah]

- (b) Katakan X_1, X_2, \dots, X_n satu sampel rawak saiz n daripada suatu taburan Poisson dengan parameter $\lambda = 1$, iaitu,

$$P(X_i = x) = \frac{e^{-1}}{x!}, \quad x = 0, 1, 2, \dots, i = 1, 2, \dots, n$$

- (i) Terbitkan f.p.m bagi X_i .
- (ii) Biarkan $\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}$
Pertimbangkan $Y_n = \sqrt{n}(\bar{X}_n - 1)$. Tunjukkan f.p.m bagi Y_n diberi oleh $\exp[-t\sqrt{n} + n(e^{t/\sqrt{n}} - 1)]$.
- (iii) Demikian, dapatkan taburan penghad bagi Y_n apabila $n \rightarrow \infty$.

[50 markah]

2. (a) Let $Y_1 < Y_2 < Y_3 < Y_4$ be the order statistics from a random sample of size 4 taken from a distribution with p.d.f.

$$f(x) = \begin{cases} 2x, & 0 < x < 1 \\ 0, & \text{elsewhere} \end{cases}$$

- (i) Obtain the joint p.d.f. of Y_3 and Y_4 .
(ii) Find the conditional p.d.f. of Y_3 given $Y_4 = y_4$
(iii) Evaluate $E(Y_3 | y_4)$.

[30 marks]

- (b) State and prove Chebychev's inequality.

[30 marks]

- (c) Let X_1, X_2, \dots, X_n be a random sample from a distribution with p.d.f.

$$f(x; \theta) = \begin{cases} e^{-(x-\theta)} \exp(-e^{-(x-\theta)}), & x \geq \theta, -\infty < \theta < \infty \\ 0, & x < \theta \end{cases}$$

- (i) Find the MLE for θ .
(ii) Find the MLE for $P_\theta\{X_1 - \theta \geq 1\}$.

[40 marks]

2. (a) Katakan $Y_1 < Y_2 < Y_3 < Y_4$ statistik tertib daripada suatu sampel rawak saiz $n = 4$ yang diambil daripada suatu taburan dengan f.k.k.

$$f(x, \theta) = \begin{cases} 2x, & 0 < x < 1 \\ 0, & \text{di tempat lain} \end{cases}$$

- (i) Dapatkan f.k.k. tercantum bagi Y_3 dan Y_4 .
(ii) Cari f.k.k. bersyarat bagi Y_3 diberi $Y_4 = y_4$.
(iii) Nilaikan $E(Y_3 | y_4)$.

[30 markah]

(b) Nyatakan dan buktikan ketaksamaan Chebychev.

[30 markah]

(c) Katakan X_1, X_2, \dots, X_n adalah suatu sampel rawak dari fungsi ketumpatan kebarangkalian

$$f(x, \theta) = \begin{cases} e^{-(x-\theta)} \exp(-e^{-(x-\theta)}), & x \geq \theta, -\infty < \theta < \infty \\ 0 & , x < \theta \end{cases}$$

(i) Carikan MLE θ .

(ii) Carikan MLE bagi $P_\theta \{X_1 - \theta \geq 1\}$.

[40 markah]

3. (a) What are the roles of the following theorems in estimation theory:

(i) Neyman – Fisher Factorization Theorem

(ii) Rao – Blackwell Theorem.

(iii) Lehmann – Scheffe Theorem.

(iv) Cramer – Rao Theorem?

Give examples to illustrate your answer.

[40 marks]

(b) Let X_1, X_2, \dots, X_n be a random sample from a distribution with p.d.f.

$$f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x}, & x > 0, \lambda > 0 \\ 0 & , \text{ elsewhere} \end{cases}$$

(i) Show that $T = \sum_{i=1}^n X_i$ is a statistic which is sufficient and complete for λ .

(ii) Let $U(X_1) = \begin{cases} 0 & \text{if } X_1 < K, \\ 1 & \text{if } X_1 \geq K, \end{cases}$

where K is a known constant. Show that $U(X_1)$ is unbiased for $e^{-K\lambda}$.

(iii) Hence, show that the uniformly minimum variance unbiased estimator for $e^{-K\lambda}$ is given by

$$g(T) = \begin{cases} \left(\frac{T-K}{T}\right)^{n-1} & \text{when } T \geq K \\ 0 & \text{when } T < K \end{cases}$$

[60 marks]

3. (a) Apakah peranan teorem-teorem berikut di dalam teori penganggaran:

(i) Teorem Pemfaktoran Neyman-Fisher,

(ii) Teorem Rao-Blackwell,

(iii) Teorem Lehmann-Scheffe,

(iv) Teorem Cramer-Rao?

Beri contoh-contoh untuk mengilustrasi jawapan anda.

[40 markah]

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- (b) Katakan X_1, X_2, \dots, X_n suatu sampel rawak daripada suatu taburan dengan f.k.k.

$$f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x}, & x > 0, \lambda > 0 \\ 0, & \text{di tempat lain} \end{cases}$$

(i) Tunjukkan $T = \sum_{i=1}^n X_i$ ialah suatu statistik yang cukup dan lengkap bagi λ .

(ii) Katakan $U(X_1) = \begin{cases} 0 & \text{jika } X_1 < K, \\ 1 & \text{jika } X_1 \geq K, \end{cases}$ di mana K ialah suatu pemalar yang diketahui. Tunjukkan $U(X_1)$ saksama bagi $e^{-K\lambda}$.

(iii) Demikian, tunjukkan bahawa penganggar saksama bervarians minimum secara seragam bagi $e^{-K\lambda}$ diberi oleh

$$g(T) = \begin{cases} \left(\frac{T-K}{T}\right)^{n-1} & \text{apabila } T \geq K \\ 0 & \text{apabila } T < K \end{cases}$$

[60 markah]

4. (a) Let X_1, X_2, \dots, X_n be a random sample from a distribution with p.d.f.

$$f(x|\theta) = \begin{cases} \theta x^{\theta-1}, & 0 < x < 1, \theta > 0 \\ 0, & \text{elsewhere} \end{cases}$$

Suppose than the prior distribution for Θ is given by the p.d.f.

$$\pi(\theta) = \begin{cases} \frac{\lambda^r \theta^{r-1} e^{-\theta\lambda}}{\Gamma(r)}, & \theta > 0 \\ 0, & \text{elsewhere} \end{cases}$$

where r and λ are known positive constants.

- (i) Find the posterior p.d.f. for Θ
 (ii) Find the Bayes estimator for θ with respect to the prior $\pi(\theta)$ above using the squared error loss function

[40 marks]

- (b) Let X_1, X_2, \dots, X_n be a random sample of size from a distribution with p.d.f.

$$f(x;\theta) = \begin{cases} \theta(1+x)^{-(1+\theta)}, & 0 < x < \infty, \theta > 0 \\ 0, & \text{elsewhere} \end{cases}$$

- (i) Find the m.l.e. for (a) θ and (b) $\frac{1}{\theta}$.
 (ii) Does there exist a complete and sufficient statistic for θ ?

- (iii) Find a uniformly minimum variance unbiased estimator for (a) $\frac{1}{\theta}$ dan (b) θ .
- (iv) Find the Cramer-Rao lower bound for the variances of unbiased estimators of $\frac{1}{\theta}$.

[60 marks]

4. (a) Katakan X_1, X_2, \dots, X_n suatu sampel rawak daripada suatu taburan dengan f.k.k.

$$f(x|\theta) = \begin{cases} \theta x^{\theta-1}, & 0 < x < 1, \theta > 0 \\ 0, & \text{di tempat lain} \end{cases}$$

Andaikan bahawa taburan priori bagi Θ diberi oleh f.k.k.

$$\pi(\theta) = \begin{cases} \frac{\lambda^r \theta^{r-1} e^{-\theta\lambda}}{\Gamma(r)}, & \theta > 0 \\ 0, & \text{di tempat lain} \end{cases}$$

di mana r and λ adalah pemalar positif yang diketahui.

- (i) Cari f.k.k. posterior bagi Θ
 (ii) Cari penganggar Bayes bagi θ terhadap prior $\pi(\theta)$ di atas dengan menggunakan fungsi kerugian ralat kuasa dua.

[40 markah]

- (b) Katakan X_1, X_2, \dots, X_n suatu sampel rawak saiz n daripada suatu taburan dengan f.k.k.

$$f(x;\theta) = \begin{cases} \theta(1+x)^{-(1+\theta)}, & 0 < x < \infty, \theta > 0 \\ 0, & \text{di tempat lain} \end{cases}$$

- (i) Cari p.k.m bagi (a) θ dan (b) $\frac{1}{\theta}$.
 (ii) Adakah wujudnya suatu statistik yang cukup dan lengkap bagi θ ?
 (iii) Cari suatu penganggar saksama bervarians minimum secara seragam bagi (a) $\frac{1}{\theta}$ dan (b) θ .
 (iv) Cari batas bawah Cramer-Rao bagi varians penganggar saksama untuk $\frac{1}{\theta}$.

[60 markah]

5. (a) State the Neyman-Pearson fundamental lemma.

[10 marks]

- (b) Describe the concepts which are relevant in the theory of hypothesis testing.

[20 marks]

- (c) Let X_1, X_2 be a random sample of size 2 from a distribution with p.d.f.

$$f(x; \theta) = \begin{cases} (1+\theta)x^\theta, & 0 \leq x \leq 1, \theta > 0 \\ 0, & \text{elsewhere} \end{cases}$$

Show that the uniformly most powerful test for testing $H_0 : \theta = 2$ versus $H_1 : \theta = \theta_1 < 2$ is given by

$$\phi(x_1, x_2) = \begin{cases} 1 & \text{if } x_1 x_2 \leq \frac{3}{4} \\ 0 & \text{if } x_1 x_2 > \frac{3}{4} \end{cases}.$$

What is the size of the above test?

[30 marks]

- (d) Let X_1, X_2, \dots, X_n be a random sample from a distribution with p.d.f. $N(\mu, \sigma^2)$, $-\infty < \mu < \infty$, $0 < \sigma^2 < \infty$.

Let Y_1, Y_2, \dots, Y_m be a random sample from a distribution with p.d.f.

$N(\eta, \tau^2)$, $-\infty < \eta < \infty$, $0 < \tau^2 < \infty$.

Suppose that the two samples are independent. Show that the likelihood ratio test for testing $H_0 : \sigma^2 = \tau^2, \mu$ and η unknown, versus $H_1 : \sigma^2 \neq \tau^2, \mu$ and η unknown, can be based on the statistic

$$F = \frac{\sum_{i=1}^n (X_i - \bar{X})^2 / (n-1)}{\sum_{j=1}^m (Y_j - \bar{Y})^2 / (m-1)},$$

where $\bar{X} = \sum_{i=1}^n X_i / n$ and $\bar{Y} = \sum_{j=1}^m Y_j / m$.

[40 marks]

5. (a) Nyatakan lema asasi Neyman-Pearson.

[10 markah]

- (b) Huraikan konsep-konsep yang berkaitan di dalam teori pengujian hipotesis.

[20 markah]

- (c) Katakan X_1, X_2 sampel rawak saiz 2 daripada suatu taburan dengan f.k.k.

$$f(x; \theta) = \begin{cases} (1+\theta)x^\theta, & 0 \leq x \leq 1, \theta > 0 \\ 0, & \text{di tempat lain} \end{cases}$$

Tunjukkan bahawa ujian yang paling berkuasa secara seragam untuk menguji $H_0 : \theta = 2$ lawan $H_1 : \theta = \theta_1 < 2$ diberi oleh

$$\phi(x_1, x_2) = \begin{cases} 1 & \text{jika } x_1 x_2 \leq \frac{3}{4} \\ 0 & \text{jika } x_1 x_2 > \frac{3}{4} \end{cases}.$$

Apakah saiz ujian di atas?

[30 markah]

(d) Katakan X_1, X_2, \dots, X_n sampel rawak daripada taburan dengan f.k.k.

$$N(\mu, \sigma^2), -\infty < \mu < \infty, 0 < \sigma^2 < \infty.$$

Katakan Y_1, Y_2, \dots, Y_m sampel rawak daripada taburan dengan f.k.k.

$$N(\eta, \tau^2), -\infty < \eta < \infty, 0 < \tau^2 < \infty.$$

Andaikan dua sampel tersebut tak bersandar. Tunjukkan bahawa ujian nisbah kebolehjadian untuk menguji $H_0 : \sigma^2 = \tau^2, \mu$ dan η anu, lawan $H_1 : \sigma^2 \neq \tau^2, \mu$ dan η anu, boleh didasarkan pada statistik.

$$F = \frac{\sum_{i=1}^n (X_i - \bar{X})^2 / (n-1)}{\sum_{j=1}^m (Y_j - \bar{Y})^2 / (m-1)},$$

$$\text{di mana } \bar{X} = \sum_{i=1}^n X_i / n \text{ dan } \bar{Y} = \sum_{j=1}^m Y_j / m.$$

[40 markah]

Taburan	Fungsi Keluampaian f	$H(x)$ $\mu = E[X]$	Varians $\sigma^2 = E[(X - \mu)^2]$	Fungsi penjanaan Moment
Seragam Diskrit	$A(x) = \frac{1}{N} I_{\{1, 2, \dots, N\}}(x)$	$\frac{N+1}{2}$	$\frac{N^2-1}{12}$	$\sum_{j=1}^N \frac{1}{N} j^t$
Bernoulli	$f(x) = p^x q^{1-x} I_{\{0, 1\}}(x)$	p	pq	$q + p e^t$
Binormal	$f(x) = \binom{n}{x} p^x q^{n-x} I_{\{0, 1, \dots, n\}}(x)$	$n p$	$n p q$	$(q + p e^t)^n$
Geometri	$f(x) = p q^x I_{\{0, 1, \dots\}}(x)$	q	$\frac{q}{p}$	$\frac{p}{1-q e^t}$
Poisson	$f(x) = e^{-\lambda} \frac{\lambda^x}{x!} I_{\{0, 1, \dots\}}(x)$	λ	λ	$e^{\lambda(e^t - 1)}$
Seragam	$f(x) = \frac{1}{b-a} I_{[a,b)}(x)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{b e^{-at} - a e^{-bt}}{(b-a)t}$
Normal	$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\{-\frac{(x-\mu)^2}{2\sigma^2}\} I_{(-\infty, \infty)}(x)$	μ	σ^2	$\exp\{\mu + \frac{1}{2}\sigma^2 t^2\}$
Eksponen	$f(x) = \lambda e^{-\lambda x} I_{[0, \infty)}(x)$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$\frac{1}{\lambda-t}, \quad t < \lambda$
Gamma	$f(x) = \frac{\lambda^n}{\Gamma(n)} e^{-\lambda x} x^{n-1} I_{[0, \infty)}(x)$	$\frac{n}{\lambda}$	$\frac{n}{\lambda^2}$	$\left(\frac{\lambda}{\lambda-t}\right)^n, \quad t < \lambda$
Khi kuasa dua	$f(x) = \left(\frac{1}{2}\right)^r 2^{\frac{1}{2}} \frac{1}{\Gamma(r/2)} e^{-x/2} x^{r/2-1} I_{[0, \infty)}(x)$	r	$2r$	$\left(\frac{1}{1-2t}\right)^{r/2}, \quad t < \frac{1}{2}$
Beta	$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} I_{(0,1)}(x)$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta)(\alpha+\beta+1)^2}$	—

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