
UNIVERSITI SAINS MALAYSIA

Peperiksaan Kursus Semasa Cuti Panjang
Sidang Akademik 2007/2008

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MAT 363 – Statistical Inference
[Pentaabiran Statistik]

Duration : 3 hours
[Masa : 3 jam]

Please check that this examination paper consists of EIGHT pages of printed material before you begin the examination.

[Sila pastikan bahawa kertas peperiksaan ini mengandungi LAPAN muka surat yang bercetak sebelum anda memulakan peperiksaan ini.]

Instructions: Answer **all four** [4] questions.

Arahan: Jawab **semua empat** [4] soalan.]

1. (a) Let $f(x) = \frac{1}{3}$, $-1 < x < 2$, zero elsewhere, be the probability density function of X . Find the cumulative distribution and the probability density function of $Y = X^2$.
(Hint: Consider $P(X^2 \leq y)$ for two cases $0 \leq y < 1$ and $1 \leq y < 4$)
[20 marks]
- (b) Let $f(x, y) = e^{-x-y}$, $0 < x < \infty$, $0 < y < \infty$, zero elsewhere, be the joint probability density function of X and Y . Then if $Z = X + Y$, find $P(Z \leq z)$, for $0 < z < \infty$. What is the probability density function of Z ?
[30 marks]
- (c) Find the variance of Y if the moment generating function is $M_Y(t) = e^{2t} / (1 - t^2)$.
[20 marks]
- (d) Let (X, Y) be continuous random variables with joint probability density function
- $$f(x, y) = \begin{cases} \frac{1}{2}, & 0 < x < y < 2 \\ 0, & \text{elsewhere} \end{cases}$$

Find the conditional probability density of X given $Y = y$.

[30 marks]

2. (a) Cast a die a number of independent times until a six appear on the up side of the die.
- (i) Find the probability mass function $p(x)$ of X , the number of casts needed to obtain the first six.
- (ii) Show that $\sum_{x=1}^{\infty} p(x) = 1$.
- (iii) Determine $P(X = 1, 3, 5, 7, \dots)$
[30 marks]
- (b) If continuous random variables (X, Y) have joint density function $f(x, y) = 2$, $0 < x < y < 1$, and zero elsewhere, find the correlation coefficient for X and Y .
[20 marks]
- (c) If X_1, X_2, \dots, X_n is a random sample from the $N(\mu, 4\sigma^2)$ distribution, and \bar{X}_m is defined as
- $$\bar{X}_m = \frac{1}{m} \sum_{i=1}^m X_i, \quad m \leq n,$$
- find the distribution for each of the following statistics:

1. (a) Biarkan $f(x) = \frac{1}{3}$, $-1 < x < 2$, sifar di tempat lain, adalah fungsi ketumpatan kebarangkalian bagi X . Cari taburan longgokan dan fungsi ketumpatan kebarangkalian bagi $Y = X^2$.
(Petunjuk: Pertimbangkan $P(X^2 \leq y)$ bagi dua kes $0 \leq y < 1$ dan $1 \leq y < 4$)
[20 markah]

- (b) Biarkan $f(x, y) = e^{-x-y}$, $0 < x < \infty$, $0 < y < \infty$, sifar di tempat lain, adalah fungsi ketumpatan kebarangkalian tercantum bagi X dan Y . Kemudian jika $Z = X + Y$, cari $P(Z \leq z)$, bagi $0 < z < \infty$. Apakah fungsi ketumpatan kebarangkalian bagi Z ?
[30 markah]

- (c) Cari varians bagi Y jika fungsi penjana momen adalah $M_Y(t) = e^{2t} / (1 - t^2)$.
[20 markah]

- (d) Biarkan (X, Y) sebagai pembolehubah rawak selanjar dengan fungsi ketumpatan kebarangkalian tercantum

$$f(x, y) = \begin{cases} \frac{1}{2}, & 0 < x < y < 2 \\ 0, & \text{di tempat lain} \end{cases}$$

Cari ketumpatan kebarangkalian bersyarat bagi X diberi $Y = y$.

[30 markah]

2. (a) Lambung sebiji dadu sebilangan kali yang tak bersandar sehingga satu nombor enam muncul pada sebelah atas dadu tersebut.

(i) Cari fungsi jisim kebarangkalian, $p(x)$ bagi X , bilangan kali yang diperlukan untuk memperoleh nombor enam yang pertama.

(ii) Tunjukkan bahawa $\sum_{x=1}^{\infty} p(x) = 1$.

(iii) Tentukan $P(X = 1, 3, 5, 7, \dots)$

[30 markah]

- (b) Jika pembolehubah rawak selanjar (X, Y) mempunyai fungsi kebarangkalian tercantum $f(x, y) = 2$, $0 < x < y < 1$, dan sifar di tempat lain, cari pekali korelasi bagi X dan Y .

[20 markah]

- (c) Jika X_1, X_2, \dots, X_n merupakan sampel rawak daripada taburan $N(\mu, 4\sigma^2)$, dan \bar{X}_m ditakrifkan sebagai,

$$\bar{X}_m = \frac{1}{m} \sum_{i=1}^m X_i, \quad m \leq n,$$

cari taburan bagi setiap statistik berikut:

- (i) $m\bar{X}_m - n\bar{X}_n$
 (ii) $(m-1)\bar{X}_m + (n-1)\bar{X}_n$
 (iii) $\frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$
 (iv) $\sum_{i=1}^m \frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$

[30 marks]

- (d) Let X_1, X_2, \dots, X_n represent a random sample from a $\text{Be}(\theta)$ distribution, $0 < \theta < 1$, and \bar{X}_n represents the corresponding sample mean. Show that $\bar{Y}_n = 1 - \bar{X}_n \xrightarrow{P} 1 - \theta$.

[20 marks]

3. (a) Let X_1, X_2, \dots, X_n represent a random sample from a Bernoulli distribution with parameter λ , $0 \leq \lambda \leq 1$. Find the Cramer-Rao lower bound for the variance of unbiased estimators of $\lambda(1 - \lambda)$.

[20 marks]

- (b) Assume that X_1, X_2, \dots, X_n is a random sample of size n that has a distribution with density function

$$f(x; \theta) = \begin{cases} (\theta + 1)x^\theta, & 0 < x < 1 \\ 0, & \text{elsewhere} \end{cases}$$

- (i) Is this distribution a member of the exponential family? Explain.
 (ii) Find a uniformly minimum variance of unbiased estimator (UMVUE) for $\theta + 1$.

[40 marks]

- (c) If X_1, X_2, \dots, X_n is a random sample from the gamma, $G(3, \alpha)$ distribution and the sample mean, $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, find

- (i) an approximate $100\gamma\%$ confidence interval for α when n is large, where $0 < \gamma < 1$.
 (ii) a $100\gamma\%$ confidence interval for α when n is small.

[40 marks]

- (i) $m\bar{X}_m - n\bar{X}_n$
(ii) $(m-1)\bar{X}_m + (n-1)\bar{X}_n$
(iii) $\frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$
(iv) $\sum_{i=1}^m \frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$

[30 markah]

- (d) Biarkan X_1, X_2, \dots, X_n mewakili sampel rawak daripada taburan $Be(\theta)$, $0 < \theta < 1$, dan \bar{X}_n mewakili min sampel yang sepadan. Tunjukkan bahawa $\bar{Y}_n = 1 - \bar{X}_n \xrightarrow{P} 1 - \theta$.

[20 markah]

3. (a) Biarkan X_1, X_2, \dots, X_n menandakan suatu sampel rawak daripada taburan Bernoulli dengan parameter λ , $0 \leq \lambda \leq 1$. Cari batas bawah Cramer-Rao bagi varians penganggar-penganggar saksama $\lambda(1 - \lambda)$.

[20 markah]

- (b) Andaikan X_1, X_2, \dots, X_n ialah suatu sampel rawak bersaiz n yang mempunyai taburan dengan fungsi ketumpatan

$$f(x; \theta) = \begin{cases} (\theta + 1)x^\theta, & 0 < x < 1 \\ 0, & \text{di tempat lain} \end{cases}$$

- (i) Adakah taburan ini ahli bagi famili eksponen? Jelaskan.
(ii) Cari penganggar saksama bervarians minimum secara seragam (PSVMS) bagi $\theta + 1$.

[40 markah]

- (c) Jika X_1, X_2, \dots, X_n suatu sampel rawak daripada taburan gama, $G(3, \alpha)$ dan min sampel, $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, cari

- (i) suatu selang keyakinan hampiran $100\gamma\%$ bagi α apabila n besar, yang mana $0 < \gamma < 1$.
(ii) suatu selang keyakinan $100\gamma\%$ bagi α apabila n kecil.

[40 markah]

4. (a) Assume that X is a single observation from a distribution having density function

$$f(x; \theta) = \theta x^{\theta-1} I_{(0,1)}(x); \quad \theta > 0.$$

For testing $H_0 : \theta \leq 1$ vs. $H_1 : \theta > 1$, the test that is used is reject H_0 if and only if $X \geq \frac{1}{2}$. Find the power function and the size of the test.

[30 marks]

- (b) Let X_1, X_2, \dots, X_n represents a random sample of size n from a distribution with density function $f(x; \alpha) = \alpha^2 x e^{-\alpha x} I_{(0,\infty)}(x)$.

- (i) Find the uniformly most powerful critical region for testing $H_0 : \alpha = 1$ vs. $H_1 : \alpha > 1$.

- (ii) For testing $H_0 : \alpha = 1$ vs. $H_1 : \alpha \neq 1$, the following test is used: Reject H_0 if and only if $|\bar{X} - 2| \geq k$. Find an approximate value of k so that the size of the Type-I error is 0.05. Assume that n is sufficiently large so that the central limit theorem can be used to find an approximate value of k .

[40 marks]

- (c) Assume that the single observation X follows a binomial, $b(n, \theta)$ distribution with probability mass function

$$f(x; \theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}, \quad x = 0, 1, 2, \dots, n.$$

Find the generalized likelihood ratio test for testing $H_0 : \theta \leq \theta_0$ vs. $H_1 : \theta > \theta_0$.

[30 marks]

4. (a) Andaikan X cerapan tunggal daripada taburan yang mempunyai fungsi ketumpatan

$$f(x; \theta) = \theta x^{\theta-1} I_{(0,1)}(x); \quad \theta > 0.$$

Untuk menguji $H_0 : \theta \leq 1$ lawan $H_1 : \theta > 1$, ujian yang digunakan ialah tolak H_0 jika dan hanya jika $X \geq \frac{1}{2}$. Cari fungsi kuasa dan saiz ujian tersebut.

[30 markah]

- (b) Biarkan X_1, X_2, \dots, X_n mewakili suatu sampel rawak saiz n daripada taburan dengan fungsi ketumpatan $f(x; \alpha) = \alpha^2 x e^{-\alpha x} I_{(0,\infty)}(x)$.

- (i) Cari rantau genting paling berkuasa secara seragam bagi menguji $H_0 : \alpha = 1$ lawan $H_1 : \alpha > 1$.

- (ii) Untuk menguji $H_0 : \alpha = 1$ lawan $H_1 : \alpha \neq 1$, ujian berikut digunakan: Tolak H_0 jika dan hanya jika $|\bar{X} - 2| \geq k$. Cari nilai hampiran k supaya saiz ralat Jenis-I ialah 0.05. Andaikan bahawa n adalah cukup besar supaya teorem had memusat boleh digunakan untuk mencari nilai hampiran k .

[40 markah]

- (c) Andaikan cerapan tunggal, X bertaburan binomial, $b(n, \theta)$ dengan fungsi jisim kebarangkalian

$$f(x; \theta) = \binom{n}{x} \theta^x (1 - \theta)^{n-x}, \quad x = 0, 1, 2, \dots, n.$$

Cari ujian nisbah kebolehdajian teritlak bagi menguji $H_0 : \theta \leq \theta_0$ lawan $H_1 : \theta > \theta_0$.

[30 markah]

APPENDIX / LAMPIRAN

Taburan	Fungsi Ketumpalan	Min	Varians	Fungsi Panjang Momen
Seragam Diskrit	$f(x) = \frac{1}{N} I_{\{1, 2, \dots, N\}}(x)$	$\frac{N+1}{2}$	$\frac{N^2 - 1}{12}$	$\sum_{j=1}^N \frac{1}{N} e^{j\mu}$
Bernoulli	$f(x) = p^x q^{1-x} I_{\{0,1\}}(x)$	p	pq	$q + pe'$
Binomial	$f(x) = \binom{n}{x} p^x q^{n-x} I_{\{0,1,\dots,n\}}(x)$	np	npq	$(q + pe')^n$
Geometri	$f(x) = pq^x I_{\{0,1,\dots\}}(x)$	$\frac{q}{p}$	$\frac{q}{p^2}$	$\frac{p}{1 - qe'}, qe' < 1$
Poisson	$f(x) = e^{-\lambda} \frac{\lambda^x}{x!} I_{\{0,1,\dots\}}(x)$	λ	λ	$\exp\{\lambda(e' - 1)\}$
Seragam	$f(x) = \frac{1}{b-a} I_{[a,b]}(x)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{e^{b\mu} - e^{a\mu}}{(b-a)\mu}, \mu \neq 0$
Normal	$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} I_{(-\infty,\infty)}(x)$	μ	σ^2	$\exp\{\mu\mu + (\sigma\mu)^2/2\}$
Eksponen	$f(x) = \lambda e^{-\lambda x} I_{(0,\infty)}(x)$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$\frac{\lambda}{\lambda - t}, t < \lambda$
Gamma	$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda x} x^{\alpha-1} I_{(0,\infty)}(x)$	$\frac{\alpha}{\lambda}$	$\frac{\alpha}{\lambda^2}$	$\left(\frac{\lambda}{\lambda - t}\right)^\alpha, t < \lambda$
Khi Kuasa Dua	$f(x) = \left(\frac{1}{2}\right)^{r/2} \frac{1}{\Gamma(r/2)} e^{-x/2} x^{(r/2)-1} I_{(0,\infty)}(x)$	r	$2r$	$\left(\frac{1}{1-2t}\right)^{r/2}, t < \frac{1}{2}$
Beta	$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} I_{(0,1)}(x)$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$	