
UNIVERSITI SAINS MALAYSIA

First Semester Examination
2012/2013 Academic Session

January 2013

MST 561 – Statistical Inference
[Pentaabiran Statistik]

Duration : 3 hours
[Masa : 3 jam]

Please check that this examination paper consists of EIGHT pages of printed material before you begin the examination.

[Sila pastikan bahawa kertas peperiksaan ini mengandungi LAPAN muka surat yang bercetak sebelum anda memulakan peperiksaan ini].

Instructions: Answer **all five** [5] questions.

[Arahan: Jawab **semua lima** [5] soalan.]

In the event of any discrepancies, the English version shall be used.

[Sekiranya terdapat sebarang percanggahan pada soalan peperiksaan, versi Bahasa Inggeris hendaklah diguna pakai].

1. (a) Let Y be a continuous random variable with $f_Y(y) = ye^{-y}, 0 \leq y$. Show that the moment generating function $M_Y(t) = \frac{1}{(1-t)^2}$. Hence, find the mean and standard deviation of the random variable Y .

[35 marks]

- (b) Let X and Y have joint probability density function $f_{X,Y}(x, y) = \begin{cases} 10xy^2, & 0 < x < y < 1 \\ 0, & \text{elsewhere} \end{cases}$
 Suppose $A = X/Y$ and $B = Y$. Find the joint probability density function of (A, B) and hence the marginal probability density of A and B .

[35 marks]

- (c) Let X have the probability density function $f(x) = 4x^3, 0 < x < 1$, zero elsewhere. Find the cumulative distribution function and the probability density function of $Y = -2\ln X^4$.

[30 marks]

1. (a) *Andaikan Y suatu pembolehubah rawak selanjar dengan $f_Y(y) = ye^{-y}, 0 \leq y$. Tunjukkan bahawa fungsi penjana momen $M_Y(t) = \frac{1}{(1-t)^2}$.*

Seterusnya, cari min dan sisihan piawai bagi pembolehubah rawak Y .

[35 markah]

- (b) *Biarkan X dan Y mempunyai fungsi ketumpatan kebarangkalian tercantum $f_{X,Y}(x, y) = \begin{cases} 10xy^2, & 0 < x < y < 1 \\ 0, & \text{selainnya} \end{cases}$
 Andaikan $A = X/Y$ dan $B = Y$. Cari fungsi ketumpatan kebarangkalian tercantum (A, B) dan seterusnya fungsi ketumpatan kebarangkalian sut bagi A dan B .*

[35 markah]

- (c) *Biarkan X mempunyai fungsi ketumpatan kebarangkalian $f(x) = 4x^3, 0 < x < 1$, sifar selainnya. Cari fungsi taburan longgokan dan fungsi ketumpatan kebarangkalian bagi $Y = -2\ln X^4$.*

[30 markah]

2. (a) (i) If the random variable X has a standard normal distribution, show using the distribution function technique that $Y = X^2$ has a chi-square distribution with degree of freedom one.
- (ii) If Y_1, Y_2, \dots, Y_n are random variables from chi-square distribution with degree of freedom one, find the distribution of $Z = \sum_{i=1}^n Y_i$ using the moment generating function approach.

[40 marks]

- (b) A random sample of size $n = 6$ is taken from the pdf $f_Y(y) = 3y^2, 0 \leq y \leq 1$. Find the probability of the fifth order statistics $P(Y_5' > 0.75)$.

[30 marks]

- (c) Let Y_1, Y_2, \dots, Y_n be a random sample from the exponential pdf $f_Y(y) = e^{-y}, y > 0$. What is the smallest n for which $P(Y_{\min} < 0.2) > 0.9$?

[30 marks]

2. (a) (i) Jika pembolehubah rawak X mempunyai taburan normal piawai, tunjukkan dengan menggunakan teknik fungsi taburan bahawa $Y = X^2$ mempunyai suatu taburan chi-kuasadua dengan darjah kebebasan satu.
- (ii) Jika Y_1, Y_2, \dots, Y_n adalah pembolehubah rawak dari taburan khi-kuasadua dengan darjah kebebasan satu, cari taburan bagi $Z = \sum_{i=1}^n Y_i$ dengan menggunakan kaedah fungsi penjana momen.

[40 markah]

- (b) Suatu sampel rawak bersaiz $n = 6$ diambil dari fungsi ketumpatan kebarangkalian $f_Y(y) = 3y^2, 0 \leq y \leq 1$. Cari kebarangkalian bagi statistik tertib kelima. $P(Y_5' > 0.75)$.

[30 markah]

- (c) Biarkan Y_1, Y_2, \dots, Y_n sebagai suatu sampel rawak dari taburan eksponen dengan fungsi ketumpatan kebarangkalian $f_Y(y) = e^{-y}, y > 0$. Apakah nilai terkecil n dengan keadaan $P(Y_{\min} < 0.2) > 0.9$?

[30 markah]

3. (a) Let W_n be a random variable with mean μ and variance $\frac{b}{n^p}$, where $p > 0$, μ and b are constants (not functions of n). Prove that W_n converges in probability to μ . [Hint: Use Chebyshev's inequality]

[20 marks]

- (b) Let \bar{X} be the mean of n independent random variables X_1, X_2, \dots, X_n from the exponential distribution with parameter θ .

- (i) Show that \bar{X} is an unbiased point estimator of $1/\theta$.
 (ii) Using the moment generating function technique, determine the distribution of \bar{X} .
 (iii) Use part (ii) to show that $Y = 2n\theta\bar{X}$ has a χ^2 distribution with $2n$ degrees of freedom.

[40 marks]

- (c) Let $X_1 = 0.23, X_2 = 0.65, X_3 = 0.10$ and $X_4 = 0.42$ be a random sample of size four from the probability density function

$$f_x(x; \theta) = \theta x^{\theta-1}, \quad 0 \leq x \leq 1.$$

Find the method of moments estimate for θ .

[40 marks]

3. (a) *Biarkan W_n sebagai suatu pembolehubah rawak dengan min μ dan varians $\frac{b}{n^p}$, yang mana $p > 0$, μ and b adalah pemalar (bukan fungsi n). Buktikan bahawa W_n menumpu secara kebarangkalian kepada μ . [Petunjuk: Guna ketaksamaan Chebyshev]*

[20 markah]

- (b) *Biarkan \bar{X} sebagai min bagi n pembolehubah rawak tak bersandar X_1, X_2, \dots, X_n dari suatu taburan eksponen dengan parameter θ .*

- (i) *Tunjukkan bahawa \bar{X} adalah suatu penganggar titik saksama bagi $1/\theta$.*
 (ii) *Guna teknik fungsi penjana momen bagi menentukan taburan bagi \bar{X} .*
 (iii) *Guna bahagian (ii) bagi menunjukkan bahawa $Y = 2n\theta\bar{X}$ mempunyai taburan χ^2 dengan darjah kebebasan $2n$.*

[40 markah]

- (c) *Biarkan $X_1 = 0.23, X_2 = 0.65, X_3 = 0.10$ dan $X_4 = 0.42$ adalah suatu sampel rawak bersaiz empat dari fungsi ketumpatan kebarangkalian*

$$f_x(x; \theta) = \theta x^{\theta-1}, \quad 0 \leq x \leq 1.$$

Cari anggaran bagi θ dengan menggunakan kaedah momen.

[40 markah]

4. (a) Let X be a random variable from the Poisson distribution with parameter θ and assume that the prior distribution of the parameter Θ has a gamma distribution with parameters α and λ which are known. Find the posterior distribution of Θ given $X = x$ and hence write the mean of the posterior distribution of Θ .

[40 marks]

- (b) Suppose the random variables Y_1, Y_2, \dots, Y_n denote the number of successes or failures (1 or 0) in each of n independent trials, where $p = P(\text{success occurs at given trial})$ is an unknown parameter such that $p_{Y_i}(k; p) = p^k(1-p)^{1-k}$, $k=0,1; 0 < p < 1$

If $Y = Y_1 + Y_2 + \dots + Y_n = \text{total number of successes}$ and define $\hat{p} = \frac{Y}{n}$.

- (i) Show that \hat{p} is an unbiased estimator for p .
- (ii) Calculate the variance for \hat{p} .
- (iii) Compute the Cramer-Rao lower bound for $p_{Y_i}(k; p)$
- (iv) Compare the results in (ii) and (iii) and discuss.

[40 marks]

- (c) Let X_1, X_2, \dots, X_n be a random sample of size n from a probability distribution with density function $f_X(x; \theta) = \theta x^{\theta-1}$, $0 \leq x \leq 1$. If a and b are two number such that $0 < a < b < 1$, construct a 100% confidence interval for θ using the pivotal quantity method.

[20 marks]

4. (a) *Biarkan X sebagai suatu pembolehubah rawak dari taburan Poisson dengan parameter θ dan andaikan taburan prior bagi parameter Θ mempunyai taburan gama dengan parameter α dan λ yang diketahui. Cari taburan posterior bagi Θ diberi $X = x$ dan seterusnya tulis min bagi taburan posterior untuk Θ .*

[40 markah]

- (b) *Andaikan pembolehubah rawak Y_1, Y_2, \dots, Y_n adalah bilangan kejayaan atau kegagalan (1 or 0) dalam setiap n percubaan yang tak bersandar, yang mana $p = P(\text{kejayaan berlaku pada suatu percubaan yang diberi})$ adalah suatu parameter yang tidak diketahui supaya $p_{Y_i}(k; p) = p^k(1-p)^{1-k}$, $k=0,1; 0 < p < 1$*

Jika $Y = Y_1 + Y_2 + \dots + Y_n = \text{jumlah bilangan kejayaan dan takrifkan}$

$$\hat{p} = \frac{Y}{n} .$$

- (i) *Tunjukkan bahawa \hat{p} suatu penganggar saksama bagi p .*
- (ii) *Kira varians bagi \hat{p} .*
- (iii) *Kira batas bawah Cramer-Rao bagi $p_{Y_i}(k; p)$*
- (iv) *Bandingkan keputusan dalam (ii) dan (iii) dan bincang.*

[40 markah]

- (c) Biarkan X_1, X_2, \dots, X_n suatu sampel rawak bersaiz n dari taburan kebarangkalian dengan fungsi ketumpatan $f_X(x; \theta) = \theta x^{\theta-1}$, $0 \leq x \leq 1$. Jika a dan b adalah dua nombor supaya $0 < a < b < 1$, bina suatu selang keyakinan $100\gamma\%$ bagi θ dengan menggunakan kaedah kuantiti pangsaan.

[20 markah]

5. (a) Suppose $n = 36$ observations are taken from a normal distribution where $\sigma = 8.0$ for the purpose of testing $H_0 : \mu = 60$ versus $H_1 : \mu \neq 60$ at the $\alpha = 0.07$ level of significance. The lead investigator skipped statistics class the day decision rules were being discussed and intends to reject H_0 if \bar{y} falls in the interval $(60 - \bar{y}^*, 60 + \bar{y}^*)$.

(i) Find \bar{y}^* .

(ii) What is the power of the test when $\mu = 62$?

[40 marks]

- (b) Let X_1, X_2, \dots, X_n be a random sample of size n from a Poisson distribution $P(\mu)$.

(i) Find the critical region of the most powerful test of size α for testing $H_0 : \mu = 1.5$ against $H_1 : \mu = 2$.

(ii) If $\alpha = 0.05$ and $n = 10$ for $H_0 : \mu = 1$ against $H_1 : \mu = 2$, determine the most powerful test.

[40 marks]

- (c) Let X_1, X_2, \dots, X_9 be a random sample of size 9 from a distribution that is $N(\mu, \sigma^2)$. If σ is known, find the length of a 95 percent confidence interval for μ if this interval is based on the random variable $\sqrt{9}(\bar{X} - \mu) / \sigma$.

[20 marks]

5. (a) Andaikan $n = 36$ cerapan diambil dari suatu taburan normal yang mana $\sigma = 8.0$ bagi tujuan menguji $H_0 : \mu = 60$ lawan $H_1 : \mu \neq 60$ pada aras keertian $\alpha = 0.07$. Penyelidik utama mengelak kelas pada hari peraturan keputusan dibincang dan ingin menolak H_0 jika \bar{y} terletak dalam selang $(60 - \bar{y}^*, 60 + \bar{y}^*)$.
- (i) Cari \bar{y}^* .
 - (ii) Apakah kuasa ujian apabila $\mu = 62$?
- [40 markah]
- (b) Biarkan X_1, X_2, \dots, X_n suatu sampel rawak bersaiz n dari taburan Poisson $P(\mu)$.
- (i) Cari rantau genting bagi ujian paling berkuasa bersaiz α bagi menguji $H_0 : \mu = 1.5$ lawan $H_1 : \mu = 2$.
 - (ii) Jika $\alpha = 0.05$ dan $n = 10$ bagi $H_0 : \mu = 1$ lawan $H_1 : \mu = 2$, tentukan ujian paling berkuasa.
- [40 markah]
- (c) Biarkan X_1, X_2, \dots, X_9 sebagai suatu sampel rawak bersaiz 9 dari suatu taburan $N(\mu, \sigma^2)$. Jika σ diketahui, cari panjang bagi selang keyakinan 95 peratus bagi μ jika selang ini adalah berdasarkan pembolehubah rawak $\sqrt{9}(\bar{X} - \mu) / \sigma$.
- [20 markah]

LAMPIRAN

| Taburan | Fungsi Ketumpatan | Min | Varians | Fungsi Penjama Momen |
|-----------------|---|-------------------------------|--|--|
| Seragam Diskrit | $f(x) = \frac{1}{N} I_{(1,2,\dots,N)}(x)$ | $\frac{N+1}{2}$ | $\frac{N^2-1}{12}$ | $\sum_{j=1}^N \frac{1}{N} e^{jt}$ |
| Bernoulli | $f(x) = p^x q^{1-x} I_{(0,1)}(x)$ | p | pq | $q + pe^t$ |
| Binomial | $f(x) = \binom{n}{x} p^x q^{n-x} I_{(0,1,\dots,n)}(x)$ | np | npq | $(q + pe^t)^n$ |
| Geometri | $f(x) = pq^x I_{(0,1,\dots)}(x)$ | $\frac{q}{p}$ | $\frac{q}{p^2}$ | $\frac{p}{1-qe^t}, qe^t < 1$ |
| Poisson | $f(x) = e^{-\lambda} \frac{\lambda^x}{x!} I_{(0,1,\dots)}(x)$ | λ | λ | $\exp\{\lambda(e^t - 1)\}$ |
| Seragam | $f(x) = \frac{1}{b-a} I_{[a,b]}(x)$ | $\frac{a+b}{2}$ | $\frac{(b-a)^2}{12}$ | $\frac{e^{bt} - e^{at}}{(b-a)t}, t \neq 0$ |
| Normal | $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\{-(x-\mu)^2 / 2\sigma^2\} I_{(-\infty,\infty)}(x)$ | μ | σ^2 | $\exp\{\mu t + (\sigma t)^2 / 2\}$ |
| Eksponen | $f(x) = \lambda e^{-\lambda x} I_{(0,\infty)}(x)$ | $\frac{1}{\lambda}$ | $\frac{1}{\lambda^2}$ | $\frac{\lambda}{\lambda-t}, t < \lambda$ |
| Gamma | $f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda x} x^{\alpha-1} I_{(0,\infty)}(x)$ | $\frac{\alpha}{\lambda}$ | $\frac{\alpha}{\lambda^2}$ | $\left(\frac{\lambda}{\lambda-t}\right)^\alpha, t < \lambda$ |
| Khi Kuasa Dua | $f(x) = \left(\frac{1}{2}\right)^{r/2} \frac{1}{\Gamma(r/2)} e^{-x/2} x^{(r/2)-1} I_{(0,\infty)}(x)$ | r | $2r$ | $\left(\frac{1}{1-2t}\right)^{r/2}, t < \frac{1}{2}$ |
| Beta | $f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} I_{(0,1)}(x)$ | $\frac{\alpha}{\alpha+\beta}$ | $\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$ | |