
UNIVERSITI SAINS MALAYSIA

Peperiksaan Kursus Semasa Cuti Panjang
Sidang Akademik 2009/2010

Jun 2010

MAT 363 – Statistical Inference
[Pentaabiran Statistik]

Duration : 3 hours
[Masa : 3 jam]

Please check that this examination paper consists of EIGHT pages of printed material before you begin the examination.

[Sila pastikan bahawa kertas peperiksaan ini mengandungi LAPAN muka surat yang bercetak sebelum anda memulakan peperiksaan ini.]

Instructions: Answer **all four** [4] questions.

Arahan: Jawab **semua empat** [4] soalan.]

In the event of any discrepancies, the English version shall be used.

[Sekiranya terdapat sebarang percanggahan pada soalan peperiksaan, versi Bahasa Inggeris hendaklah diguna pakai].

1. (a) If X has a normal distribution, $N(0, \sigma^2)$, with density function given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{x^2}{2\sigma^2}\right) I_{(-\infty, \infty)}(x),$$

find the density for $Y = |X|$.

[30 marks]

- (b) Show that if X and Y are independent positive continuous random variables, then the density function for the random variable $Z = \frac{X}{Y}$ is

$$g(z) = \int_0^{\infty} y f_X(zy) f_Y(y) dy.$$

[30 marks]

- (c) The joint density function X_1 and X_2 is

$$f(x_1, x_2) = \frac{1}{2\pi} e^{-(x_1^2 + x_2^2)/2}, \quad -\infty < x_1, x_2 < \infty.$$

Let $Y_1 = X_1 + X_2$ and $Y_2 = X_1 - X_2$. Find the joint density function for Y_1 and Y_2 .

[40 marks]

2. (a) Let $Y_1 < Y_2 < Y_3 < Y_4$ be the order statistics of a random sample of size 4 from the distribution having probability density function

$$f(x) = e^{-x}, 0 < x < \infty, \text{ zero elsewhere.}$$

Find $P(3 \leq Y_4)$.

[30 marks]

- (b) Differentiate the moment-generating function for the geometric random variable and show that $E(X) = (1-p)/p$ and $\text{Var}(X) = (1-p)/p^2$.

[30 marks]

- (c) (i) Let U has an F distribution with parameters r and s . Using the related theorem, prove that $\frac{1}{U}$ has an F distribution with parameters s and r .

- (ii) Let $T = \frac{W}{\sqrt{V/r}}$, where the independent variables W and V follow the standard normal distribution and the chi-square distribution with r degrees of freedom, respectively. Using the related theorem, show that T^2 has an F distribution with parameters 1 and r .

[40 marks]

1. (a) Jika X mempunyai taburan normal, $N(0, \sigma^2)$, dengan fungsi ketumpatan yang diberikan oleh

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{x^2}{2\sigma^2}\right) I_{(-\infty, \infty)}(x),$$

cari fungsi ketumpatan untuk $Y = |X|$.

[30 markah]

- (b) Tunjukkan bahawa jika X dan Y pembolehubah rawak selangar positif yang tak bersandar, maka fungsi ketumpatan pembolehubah rawak $Z = \frac{X}{Y}$ ialah

$$g(z) = \int_0^{\infty} y f_X(zy) f_Y(y) dy.$$

[30 markah]

- (c) Fungsi ketumpatan tercantum X_1 dan X_2 ialah

$$f(x_1, x_2) = \frac{1}{2\pi} e^{-(x_1^2 + x_2^2)/2}, \quad -\infty < x_1, x_2 < \infty.$$

Biarkan $Y_1 = X_1 + X_2$ dan $Y_2 = X_1 - X_2$. Cari fungsi ketumpatan tercantum Y_1 dan Y_2 .

[40 markah]

2. (a) Biarkan $Y_1 < Y_2 < Y_3 < Y_4$ sebagai statistik tertib bagi suatu sampel rawak saiz 4 daripada suatu taburan dengan fungsi ketumpatan kebarangkalian

$$f(x) = e^{-x}, \quad 0 < x < \infty, \text{ sifar di tempat lain.}$$

Cari $P(3 \leq Y_4)$.

[30 markah]

- (b) Bezakan fungsi penjana momen bagi pembolehubah rawak geometri dan tunjukkan bahawa $E(X) = (1-p)/p$ dan $\text{Var}(X) = (1-p)/p^2$.

[30 markah]

- (c) (i) Biarkan U mempunyai taburan F dengan parameter r dan s . Dengan menggunakan teorem yang berkaitan, buktikan bahawa $\frac{1}{U}$ mempunyai taburan F dengan parameter s dan r .

- (ii) Biarkan $T = \frac{W}{\sqrt{V/r}}$, yang mana pembolehubah tak bersandar W dan V , masing-masing mengikuti taburan normal piawai dan taburan khi kuasa dua dengan r darjah kebebasan. Dengan menggunakan teorem yang berkaitan, tunjukkan bahawa T^2 mempunyai taburan F dengan parameter 1 dan r .

[40 markah]

...4/-

3. (a) Let X_1, X_2, \dots, X_n represent a random sample from the $N(\mu, 1)$ distribution. Show that the minimum variance of an arbitrary unbiased estimator μ^2 from the Cramer-Rao's inequality is $\frac{4\mu^2}{n}$.

[30 marks]

- (b) Assume that X_1, X_2, \dots, X_n represent a random sample from a common distribution with probability density function

$$f(x; \theta) = \begin{cases} \theta(1-\theta)^{x-1}x^{\theta-1}, & 0 < x \leq 1, \frac{1}{2} < \theta < 1. \\ 0 & , \text{ elsewhere} \end{cases}$$

Find the maximum likelihood estimator for θ .

[30 marks]

- (c) Assume that X_1, X_2, \dots, X_n is a random sample from the $G(5, \theta)$ distribution. By using the sample mean $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, derive the

- (i) approximate $100\gamma\%$ ($0 < \gamma < 1$) confidence interval for θ when n is large.
 (ii) $100\gamma\%$ ($0 < \gamma < 1$) confidence interval for θ when n is small.

[40 marks]

3. (a) Biarkan X_1, X_2, \dots, X_n mewakili suatu sampel rawak daripada taburan $N(\mu, 1)$. Tunjukkan bahawa varians minimum sebarang penganggar saksama untuk μ^2 daripada ketaksamaan Cramer-Rao ialah $\frac{4\mu^2}{n}$.

[30 markah]

- (b) Andaikan bahawa X_1, X_2, \dots, X_n mewakili sampel rawak daripada taburan sepunya dengan fungsi ketumpatan kebarangkalian

$$f(x, \theta) = \begin{cases} \theta (1-\theta)^{-1} x^{2\theta-1/1-\theta}, & 0 < x \leq 1, \frac{1}{2} < \theta < 1 \\ 0 & \text{di tempat lain} \end{cases}$$

Cari penganggar kebolehdajadian maksimum bagi θ .

[30 markah]

- (c) Andaikan bahawa X_1, X_2, \dots, X_n adalah suatu sampel rawak daripada taburan $G(5, \theta)$. Dengan menggunakan min sampel $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, terbitkan

- (i) selang keyakinan hampiran $100\gamma\%$ ($0 < \gamma < 1$) bagi θ apabila n besar.
(ii) selang keyakinan $100\gamma\%$ ($0 < \gamma < 1$) bagi θ apabila n kecil.

[40 markah]

4. (a) Assume that X_1, X_2, \dots, X_n is a random sample from the normal distribution with mean, μ and an unknown variance. Then

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1)$$

is a pivotal quantity. Find the shortest 100γ percent confidence interval for μ .

[30 marks]

- (b) Assume that the random variable X has the density function

$$f(x, \beta) = \begin{cases} \beta e^{-\beta x}, & 0 < x < \infty \\ 0, & \text{elsewhere} \end{cases}$$

A random sample of size two, X_1 and X_2 , is taken to test $H_0: \beta = \frac{1}{2}$ versus $H_1: \beta = \frac{1}{4}$. The test used is based on the critical region $C = \{x_1, x_2\}; x_1 + x_2 \geq 9.5$. Find the power of the test used.

[30 marks]

- (c) Let X_1, X_2, \dots, X_n represent a random sample from the Poisson distribution with density function

$$f(x, \lambda) = e^{-\lambda} \frac{\lambda^x}{x!} I_{\{0,1,\dots\}}(x).$$

Find the uniformly most powerful (UMP) test of size $\alpha = 0.1$ to test $H_0: \lambda = 2$ versus $H_1: \lambda > 2$. Assume that the sample size is $n = 30$.

[40 marks]

4. (a) Andaikan X_1, X_2, \dots, X_n sampel rawak daripada taburan normal dengan min, μ dan varians yang tidak diketahui. Maka

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1)$$

ialah suatu kuantiti pangsaan. Cari selang keyakinan 100γ peratus tersingkat bagi μ .

[30 markah]

- (b) Andaikan pembolehubah rawak X mempunyai fungsi ketumpatan

$$f(x, \beta) = \begin{cases} \beta e^{-\beta x}, & 0 < x < \infty \\ 0, & \text{dtl} \end{cases}$$

Sampel rawak saiz dua, X_1 dan X_2 , diambil bagi menguji $H_0: \beta = \frac{1}{2}$ lawan $H_1: \beta = \frac{1}{4}$. Ujian yang digunakan adalah berdasarkan rantau genting $C = \{x_1, x_2\}; x_1 + x_2 \geq 9.5$. Cari kuasa bagi ujian yang digunakan.

[30 markah]

- (c) Biarkan X_1, X_2, \dots, X_n menandai sampel rawak daripada taburan Poisson dengan fungsi ketumpatan

$$f(x, \lambda) = e^{-\lambda} \frac{\lambda^x}{x!} I_{(0,1,\dots)}(x)$$

Cari ujian paling berkuasa secara seragam (UPBS) bersaiz $\alpha = 0.1$ bagi menguji $H_0: \lambda = 2$ lawan $H_1: \lambda > 2$. Andaikan bahawa saiz sampel ialah $n = 30$.

[40 markah]

APPENDIX / LAMPIRAN

Tabaran	Fungsi Ketumpatan	Min	Varians	Fungsi Penjanaan Momen
Seragam Diskrit	$f(x) = \frac{1}{N} I_{(0, 2, \dots, N)}(x)$	$\frac{N+1}{2}$	$\frac{N^2-1}{12}$	$\sum_{j=1}^N \frac{1}{N} e^{jt}$
Bernoulli	$f(x) = p^x q^{1-x} I_{(0,1)}(x)$	p	pq	$q + pe^t$
Binomial	$f(x) = \binom{n}{x} p^x q^{n-x} I_{(0,1,\dots,n)}(x)$	np	npq	$(q + pe^t)^n$
Geometri	$f(x) = pq^x I_{(0,1,\dots)}(x)$	$\frac{q}{p}$	$\frac{q}{p^2}$	$\frac{p}{1-qe^t}, qe^t < 1$
Poisson	$f(x) = e^{-\lambda} \frac{\lambda^x}{x!} I_{(0,1,\dots)}(x)$	λ	λ	$\exp\{\lambda(e^t - 1)\}$
Seragam	$f(x) = \frac{1}{b-a} I_{(a,b)}(x)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{e^{bt} - e^{at}}{(b-a)t}, t \neq 0$
Normal	$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\{-(x-\mu)^2 / 2\sigma^2\} I_{(-\infty,\infty)}(x)$	μ	σ^2	$\exp\{i\mu t + (\sigma t)^2 / 2\}$
Eksponen	$f(x) = \lambda e^{-\lambda x} I_{(0,\infty)}(x)$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$\frac{\lambda}{\lambda-t}, t < \lambda$
Gama	$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda x} x^{\alpha-1} I_{(0,\infty)}(x)$	$\frac{\alpha}{\lambda}$	$\frac{\alpha}{\lambda^2}$	$\left(\frac{\lambda}{\lambda-t}\right)^\alpha, t < \lambda$
Khi Kuasa Dua	$f(x) = \left(\frac{1}{2}\right)^{r/2} \frac{1}{\Gamma(r/2)} e^{-x/2} x^{(r/2)-1} I_{(0,\infty)}(x)$	r	$2r$	$\left(\frac{1}{1-2t}\right)^{r/2}, t < \frac{1}{2}$
Beta	$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} I_{(0,1)}(x)$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$	