
UNIVERSITI SAINS MALAYSIA

First Semester Examination
Academic Session 2007/2008

October/November 2007

MAT 363 – Statistical Inference
[Pentaabiran Statistik]

Duration : 3 hours
[Masa : 3 jam]

Please check that this examination paper consists of EIGHT pages of printed material before you begin the examination.

[Sila pastikan bahawa kertas peperiksaan ini mengandungi LAPAN muka surat yang bercetak sebelum anda memulakan peperiksaan ini.]

Instructions: Answer **all four** [4] questions.

[Arahan: Jawab **semua empat** [4] soalan.]

...2/-

1. (a) Let $f(x) = \frac{1}{3}$, $-1 < x < 2$, zero elsewhere, be the probability density function of X . Find the cumulative distribution function of $Y = X^2$. Hence, find the probability density function for Y .

[30 marks]

- (b) Let X have a Poisson distribution with parameter λ . Let the conditional probability mass function of Y given $X = n$ be $P(Y = y | X = n) = \binom{n}{y} p^y (1-p)^{n-y}$. Find $E(Y)$.

(Hint: Find $E(Y | X = n)$ first)

[20 marks]

- (c) Let X have the probability density function $f(x) = 4x^3$, $0 < x < 1$, zero elsewhere. Find the cumulative distribution function and the probability density function of $Y = -2 \ln X^4$.

[20 marks]

- (d) The joint discrete random variables (X, Y) have probability values as shown below:

$$P(X=0, Y=0) = \frac{4}{25} \quad P(X=0, Y=1) = \frac{6}{25} \quad P(X=1, Y=0) = \frac{6}{25}$$

$$P(X=1, Y=1) = \frac{9}{25}$$

Find $E(X|Y=0)$ and $E(X|Y=1)$

[30 marks]

2. (a) Let $Y_1 < Y_2 < Y_3 < Y_4$ be the order statistics of a random sample of size 4 from the distribution having probability density function

$$f(x) = e^{-x}, 0 < x < \infty, \text{ zero elsewhere. Find } P(3 \leq Y_4).$$

[30 marks]

- (b) Differentiate the moment-generating function for the geometric random variable and verify the expressions given for $E(X) = (1-p)/p$ and $\text{Var}(X) = (1-p)/p^2$.

[20 marks]

- (c) (i) Let U have an F distribution with parameters r and s . Using the related theorem, prove that $\frac{1}{U}$ have an F distribution with parameters s and r .

- (ii) Let $T = \frac{W}{\sqrt{V/r}}$, where the independent variables W and V follow the

standard normal distribution and the chi-square distribution with r degrees of freedom, respectively. Using the related theorem, show that T^2 has an F distribution with parameters 1 and r .

[30 marks]

...3/-

1. (a) Andaikan $f(x) = \frac{1}{3}$, $-1 < x < 2$, sifar di tempat lain, adalah fungsi ketumpatan kebarangkalian bagi X . Cari fungsi taburan longgokan bagi $Y = X^2$. Seterusnya, cari fungsi ketumpatan kebarangkalian Y .

[30 markah]

- (b) Andaikan X mempunyai taburan Poisson dengan parameter λ . Biarkan fungsi jisim kebarangkalian bersyarat bagi Y diberi $X = n$ sebagai $P(Y = y | X = n) = \binom{n}{y} p^y (1-p)^{n-y}$. Cari $E(Y)$.

(Petunjuk: Cari $E(Y | X = n)$ dahulu)

[20 markah]

- (c) Andaikan X mempunyai fungsi ketumpatan kebarangkalian $f(x) = 4x^3$, $0 < x < 1$, sifar di tempat lain. Cari fungsi taburan longgokan dan fungsi ketumpatan kebarangkalian bagi $Y = -2 \ln X^4$.

[20 markah]

- (d) Pembolehubah rawak diskret tercantum (X, Y) mempunyai nilai kebarangkalian seperti yang ditunjukkan di bawah:

$$P(X=0, Y=0) = \frac{4}{25} \quad P(X=0, Y=1) = \frac{6}{25} \quad P(X=1, Y=0) = \frac{6}{25}$$

$$P(X=1, Y=1) = \frac{9}{25}$$

Cari $E(X|Y=0)$ dan $E(X|Y=1)$

[30 markah]

2. (a) Andaikan $Y_1 < Y_2 < Y_3 < Y_4$ adalah statistik tertib bagi suatu sampel rawak saiz 4 daripada suatu taburan dengan fungsi ketumpatan kebarangkalian $f(x) = e^{-x}$, $0 < x < \infty$, sifar di tempat lain. Cari $P(3 \leq Y_4)$.

[30 markah]

- (b) Bezakan fungsi penjadi momen bagi pembolehubah rawak geometri dan tunjukkan ungkapan yang diberi bagi $E(X) = (1-p)/p$ dan $Var(X) = (1-p)/p^2$.

[20 markah]

- (c) (i) Biarkan U mempunyai taburan F dengan parameter r dan s . Dengan menggunakan teorem yang berkaitan, buktikan bahawa $\frac{1}{U}$ mempunyai taburan F dengan parameter s dan r .

- (ii) Biarkan $T = \frac{W}{\sqrt{V/r}}$, yang mana pembolehubah tak bersandar W dan V ,

masing-masing mengikuti taburan normal piawai dan taburan khi kuasa dua dengan r darjah kebebasan. Dengan menggunakan teorem yang berkaitan, tunjukkan bahawa T^2 mempunyai taburan F dengan parameter 1 dan r .

[30 markah]

...4/-

(d) Let the probability mass function of Y_n be $P(Y_n = n) = 1$. Show that Y_n does not have a limiting distribution.

[20 marks]

3. (a) Suppose a random sample, Y_1, Y_2, \dots, Y_n , is taken from a $N(\mu, \sigma^2)$ distribution, where σ^2 is known.

(i) Find the Cramer-Rao lower bound for the variance of an unbiased estimator of μ .

(ii) Is $\bar{Y} = \frac{\sum_{i=1}^n Y_i}{n}$ an efficient estimator for μ ?

[30 marks]

(b) Given that $Y_1 = 2.3$, $Y_2 = 1.9$ and $Y_3 = 4.6$ is a random sample from a distribution with probability density function (pdf)

$$f_Y(y; \theta) = \frac{y^3 e^{-y/\theta}}{6\theta^4}, \quad y \geq 0.$$

Calculate the maximum likelihood estimate for θ .

[20 marks]

(c) Let \bar{X} and \bar{Y} be the means of two independent samples, each sample of size n and having the $N(\mu_1, \sigma^2)$ and $N(\mu_2, \sigma^2)$ distributions, respectively, where the common variance, σ^2 is known. Find n so that

$$P\left(\bar{X} - \bar{Y} - \frac{\sigma}{5} < \mu_1 - \mu_2 < \bar{X} - \bar{Y} + \frac{\sigma}{5}\right) = 0.90.$$

[20 marks]

(d) (i) Define the pivotal quantity.

(ii) Assume that X_1, X_2, \dots, X_n is a random sample having a $N(5, \sigma^2)$ distribution, where σ^2 is unknown. Is the random variable $\frac{\bar{X}}{\sqrt{\sigma}}$ a pivotal quantity? Explain.

[30 marks]

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(d) Biarkan fungsi jisim kebarangkalian bagi Y_n sebagai $P(Y_n = n) = 1$. Tunjukkan bahawa Y_n tidak mempunyai taburan penghad.

[20 markah]

3. (a) Andaikan suatu sampel rawak, Y_1, Y_2, \dots, Y_n , diambil daripada taburan $N(\mu, \sigma^2)$, yang mana σ^2 diketahui.

(i) Cari batas bawah Cramer-Rao untuk varians penganggar saksama μ .

(ii) Adakah $\bar{Y} = \frac{\sum_{i=1}^n Y_i}{n}$ suatu penganggar cekap bagi μ ?

[30 markah]

(b) Diberi $Y_1 = 2.3$, $Y_2 = 1.9$ dan $Y_3 = 4.6$ suatu sampel rawak daripada taburan dengan fungsi ketumpatan kebarangkalian (fkk)

$$f_Y(y; \theta) = \frac{y^3 e^{-y/\theta}}{6\theta^4}, \quad y \geq 0.$$

Kirakan anggaran kebolehjadian maksimum bagi θ .

[20 markah]

(c) Biarkan \bar{X} dan \bar{Y} sebagai min bagi dua sampel tak bersandar, setiap sampel dengan saiz n dan masing-masing mempunyai taburan $N(\mu_1, \sigma^2)$ dan $N(\mu_2, \sigma^2)$, yang mana varians sepunya, σ^2 diketahui. Cari n supaya

$$P\left(\bar{X} - \bar{Y} - \frac{\sigma}{5} < \mu_1 - \mu_2 < \bar{X} - \bar{Y} + \frac{\sigma}{5}\right) = 0.90.$$

[20 markah]

(d) (i) Takrifkan kuantiti pangsaan.

(ii) Andaikan X_1, X_2, \dots, X_n suatu sampel rawak yang mempunyai taburan $N(5, \sigma^2)$, yang mana σ^2 tidak diketahui. Adakah pembolehubah rawak $\frac{\bar{X}}{\sqrt{\sigma}}$ suatu kuantiti pangsaan? Jelaskan.

[30 markah]

...6/-

4. (a) Let Y have a binomial distribution with parameters n and p . We reject $H_0 : p = \frac{1}{2}$ and accept $H_1 : p > \frac{1}{2}$ if $Y \geq c$. Find n and c that gives a power function $\pi(p)$, such that $\pi\left(\frac{1}{2}\right) = 0.10$ and $\pi\left(\frac{2}{3}\right) = 0.95$, approximately.

[20 marks]

- (b) If X_1, X_2, \dots, X_n is a random sample from a distribution having pdf $f(x; \theta) = \theta x^{\theta-1}$, $0 < x < 1$, zero elsewhere, show using the Neyman-Pearson Theorem that the best critical region for testing $H_0 : \theta = 1$ vs. $H_1 : \theta = 2$ is

$$C = \left\{ (x_1, x_2, \dots, x_n) : c \leq \prod_{i=1}^n x_i \right\}.$$

[20 marks]

- (c) Based on a random sample of size 10 from a $N(0, \sigma^2)$ distribution, find a uniformly most powerful (UMP) test of size $\alpha = 0.05$ for testing $H_0 : \sigma^2 = 1$ vs. $H_1 : \sigma^2 > 1$.

[30 marks]

- (d) Assume that X_1, X_2, \dots, X_n is a random sample from a $N(0, \theta)$ distribution, where $\theta > 0$. Find the likelihood ratio (LR) test for testing $H_0 : \theta = \theta_0$ vs. $H_1 : \theta > \theta_0$.

[30 marks]

...7/-

4. (a) Biarkan Y mempunyai taburan binomial dengan parameter n dan p . Kita menolak $H_0 : p = \frac{1}{2}$ dan menerima $H_1 : p > \frac{1}{2}$ jika $Y \geq c$. Cari n dan c yang memberikan fungsi kuasa $\pi(p)$, sedemikian sehingga $\pi\left(\frac{1}{2}\right) = 0.10$ dan $\pi\left(\frac{2}{3}\right) = 0.95$, secara hampiran.

[20 markah]

- (b) Jika X_1, X_2, \dots, X_n ialah sampel rawak daripada suatu taburan yang mempunyai fkk $f(x; \theta) = \theta x^{\theta-1}$, $0 < x < 1$, sifar di tempat lain, tunjukkan dengan menggunakan Teorem Neyman-Pearson bahawa rantau genting terbaik untuk menguji $H_0 : \theta = 1$ lawan $H_1 : \theta = 2$ ialah $C = \left\{ (x_1, x_2, \dots, x_n) : c \leq \prod_{i=1}^n x_i \right\}$.

[20 markah]

- (c) Berdasarkan sampel rawak saiz 10 daripada taburan $N(0, \sigma^2)$, cari ujian paling berkuasa secara seragam (UPBS) saiz $\alpha = 0.05$ untuk menguji $H_0 : \sigma^2 = 1$ lawan $H_1 : \sigma^2 > 1$.

[30 markah]

- (d) Andaikan X_1, X_2, \dots, X_n ialah sampel rawak daripada taburan $N(0, \theta)$, yang mana $\theta > 0$. Cari ujian nisbah kebolehdian (NK) bagi menguji $H_0 : \theta = \theta_0$ lawan $H_1 : \theta > \theta_0$.

[30 markah]

APPENDIX / LAMPIRAN

Taburan	Fungsi Ketumpatan	Min	Varians	Fungsi Penjara Momen
Seragam Diskrit	$f(x) = \frac{1}{N} I_{(0,1,\dots,n)}(x)$	$\frac{N+1}{2}$	$\frac{N^2-1}{12}$	$\sum_{j=1}^N \frac{1}{N} e^{j\mu}$
Bernoulli	$f(x) = p^x q^{1-x} I_{(0,1)}(x)$	p	pq	$q + pe^p$
Binomial	$f(x) = \binom{n}{x} p^x q^{n-x} I_{(0,1,\dots,n)}(x)$	np	npq	$(q + pe^p)^n$
Geometri	$f(x) = pq^x I_{(0,1,\dots)}(x)$	$\frac{q}{p}$	$\frac{q}{p^2}$	$\frac{p}{1-qe^p}, qe^p < 1$
Poisson	$f(x) = e^{-\lambda} \frac{\lambda^x}{x!} I_{(0,1,\dots)}(x)$	λ	λ	$\exp\{\lambda(e^p - 1)\}$
Seragam	$f(x) = \frac{1}{b-a} I_{(a,b)}(x)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{e^b - e^a}{(b-a)^2}, t \neq 0$
Normal	$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} I_{(-\infty,\infty)}(x)$	μ	σ^2	$\exp\{\mu t + (\sigma t)^2 / 2\}$
Eksponen	$f(x) = \lambda e^{-\lambda x} I_{(0,\infty)}(x)$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$\frac{\lambda}{\lambda-t}, t < \lambda$
Gamma	$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda x} x^{\alpha-1} I_{(0,\infty)}(x)$	$\frac{\alpha}{\lambda}$	$\frac{\alpha}{\lambda^2}$	$\left(\frac{\lambda}{\lambda-t}\right)^\alpha, t < \lambda$
Khi Kuasa Dua	$f(x) = \left(\frac{1}{2}\right)^{r/2} \frac{1}{\Gamma(r/2)} e^{-x/2} x^{(r/2)-1} I_{(0,\infty)}(x)$	r	$2r$	$\left(\frac{1}{1-2t}\right)^{r/2}, t < \frac{1}{2}$
Beta	$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} I_{(0,1)}(x)$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$	

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