
UNIVERSITI SAINS MALAYSIA

Peperiksaan Kursus Semasa Cuti Panjang
Sidang Akademik 2008/2009

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MAT 363 – Statistical Inference
[Pentaabiran Statistik]

Duration : 3 hours
[Masa: 3 jam]

Please check that this examination paper consists of TEN pages of printed material before you begin the examination.

[Sila pastikan bahawa kertas peperiksaan ini mengandungi SEPULUH muka surat yang bercetak sebelum anda memulakan peperiksaan ini].

Instructions: Answer **all four** [4] questions.

[Arahan: Jawab **semua empat** [4] soalan.]

1. (a) Find the value of c such that $f(x)$ is a probability mass function for each of the following section.

(i) $f(x) = c\left(\frac{1}{4}\right)^x$, $x = 1, 2, \dots$, and zero elsewhere.

(ii) $f(x) = (2x + 1)c$, $x = 0, 1, 2, \dots, 6$, and zero elsewhere.

(iii) $f(x) = xc$, $x = 1, 2, \dots, N$ and zero elsewhere.

[30 marks]

- (b) Assume that the random variable X is gamma distributed, that is $X \sim G(\alpha, \lambda)$.

- (i) Show that the moment generating function of X is $\left(\frac{\lambda}{\lambda - t}\right)^\alpha$, for $t < \lambda$ if given that the probability density function of X is

$$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} I_{(0, \infty)}(x), \alpha > 0, \lambda > 0$$

- (ii) From the moment generating function in (i), find the mean of X .
 (iii) Hence, find the variance of X .

[30 marks]

- (c) Let $f(x, y) = 2$, $0 < x < y < 1$ and zero elsewhere be the joint probability density function of X and Y .

- (i) Find the conditional expectation of X given $Y = y$.
 (ii) Find the conditional expectation of Y given $X = x$.

[40 marks]

1. (a) Cari nilai c supaya $f(x)$ untuk setiap bahagian yang berikut merupakan suatu fungsi jisim kebarangkalian.

(i) $f(x) = c\left(\frac{1}{4}\right)^x$, $x = 1, 2, \dots$, dan sifar sebaliknya.

(ii) $f(x) = (2x + 1)c$, $x = 0, 1, 2, \dots, 6$, dan sifar sebaliknya.

(iii) $f(x) = xc$, $x = 1, 2, \dots, N$ dan sifar sebaliknya.

[30 markah]

- (b) Andaikan pembolehubah rawak X bertaburan gama, iaitu $X \sim G(\alpha, \lambda)$.

(i) Tunjukkan bahawa fungsi penjana momen X ialah $\left(\frac{\lambda}{\lambda - t}\right)^\alpha$, bagi $t < \lambda$ jika diberi fungsi ketumpatan kebarangkalian X ialah

$$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} I_{(0, \infty)}(x), \alpha > 0, \lambda > 0$$

(ii) Daripada fungsi penjana momen di (i), cari min X .

(iii) Seterusnya, cari varians X .

[30 markah]

- (c) Biarkan $f(x,y) = 2$, $0 < x < y < 1$ dan sifar sebaliknya sebagai fungsi ketumpatan kebarangkalian tercantum X dan Y .

(i) Cari jangkaan bersyarat X diberi $Y = y$.

(ii) Cari jangkaan bersyarat Y diberi $X = x$.

[40 markah]

2. (a) If X_1, X_2, \dots, X_n is a random sample from the distribution $N(\mu, 4\sigma^2)$, and \bar{X}_m is defined as $\bar{X}_m = \frac{1}{m} \sum_{i=1}^m X_i$, $m \leq n$,

Find the distribution of the following statistics:

- (i) $m\bar{X}_m - n\bar{X}_n$
 (ii) $(m-1)\bar{X}_m + (n-1)\bar{X}_n$
 (iii) $\frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$
 (iv) $\sum_{i=1}^m \frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$

[40 marks]

- (b) If Y_1, Y_2, \dots, Y_5 are order statistics from a random sample of size 5 with density function

$$f(x) = \begin{cases} \beta e^{-\beta x}, & x \geq 0 \\ 0, & \text{elsewhere} \end{cases}$$

show that Y_3 and $Y_5 - Y_3$ are independent.

[30 marks]

- (c) For a random sample of size, n , find the method of moment estimator for θ_1 and θ_2 if the random sample is taken from the distribution $G(\theta_1, \theta_2)$.

[30 marks]

2. (a) Jika X_1, X_2, \dots, X_n merupakan sampel rawak daripada taburan $N(\mu, 4\sigma^2)$, dan \bar{X}_m ditakrifkan sebagai, $\bar{X}_m = \frac{1}{m} \sum_{i=1}^m X_i$, $m \leq n$,

Cari taburan setiap statistik berikut:

- (i) $m\bar{X}_m - n\bar{X}_n$
 (ii) $(m-1)\bar{X}_m + (n-1)\bar{X}_n$
 (iii) $\frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$
 (iv) $\sum_{i=1}^m \frac{m(\bar{X}_m - \mu)^2}{4\sigma^2}$

[40 markah]

- (b) Jika Y_1, Y_2, \dots, Y_5 ialah statistik tertib daripada suatu sampel rawak saiz 5 dengan fungsi ketumpatan

$$f(x) = \begin{cases} \beta e^{-\beta x}, & x \geq 0 \\ 0, & \text{di tempat lain} \end{cases}$$

tunjukkan bahawa Y_3 dan $Y_5 - Y_3$ adalah tak bersandar.

[30 markah]

- (c) Untuk sampel rawak saiz, n , cari penganggar kaedah momen untuk θ_1 dan θ_2 jika sampel rawak itu diambil daripada taburan $G(\theta_1, \theta_2)$.

[30 markah]

3. (a) Based on a random sample of size n , use the factorization theorem to find a sufficient statistic for each of the following distributions:

(i) $f(x; \theta) = \frac{1}{\theta} I_{(0, \theta)}(x); \theta > 0.$

(ii) $f(x; \theta) = \theta^2 x e^{-x\theta} I_{(0, \infty)}(x); \theta > 0.$

[20 marks]

- (b) Based on a random sample of size n , find a complete sufficient statistic for each of the following distributions:

(i) $f(x; \theta) = (\theta + 1)x^\theta I_{(0, 1)}(x); \theta > -1.$

(ii) $f(x; \theta) = \frac{\log \theta}{\theta - 1} \theta^x I_{(0, 1)}(x); \theta > 1.$

[20 marks]

- (c) Assume that X_1, X_2, \dots, X_n is a random sample from the exponential distribution with density function

$$f(x; \theta) = \theta e^{-\theta x} I_{(0, \infty)}(x); \theta > 0.$$

Find the uniformly minimum variance unbiased estimator (UMVUE) for the median of the distribution.

[30 marks]

- (d) Assume that X_1, X_2, \dots, X_n denotes a random sample from the $P_0(\lambda)$ distribution with density function

$$f(x; \lambda) = \frac{e^{-\lambda} \lambda^x}{x!}, \quad x = 0, 1, 2, \dots$$

Find the Cramer-Rao lower bound for the variance of the unbiased estimator of $e^{-\lambda}$.

[30 marks]

3. (a) Berdasarkan suatu sampel rawak bersaiz n , guna teorem pemfaktoran untuk mencari statistic cukup bagi setiap taburan berikut:

$$(i) \quad f(x; \theta) = \frac{1}{\theta} I_{(0, \theta)}(x); \quad \theta > 0.$$

$$(ii) \quad f(x; \theta) = \theta^2 x e^{-x\theta} I_{(0, \infty)}(x); \quad \theta > 0.$$

[20 markah]

- (b) Berdasarkan suatu sampel rawak bersaiz n , cari suatu statistik cukup dan lengkap bagi setiap yang berikut:

$$(i) \quad f(x; \theta) = (\theta + 1)x^\theta I_{(0, 1)}(x); \quad \theta > -1.$$

$$(ii) \quad f(x; \theta) = \frac{\log \theta}{\theta - 1} \theta^x I_{(0, 1)}(x); \quad \theta > 1.$$

[20 markah]

- (c) Andaikan X_1, X_2, \dots, X_n sampel rawak daripada taburan eksponen yang mempunyai fungsi ketumpatan

$$f(x; \theta) = \theta e^{-\theta x} I_{(0, \infty)}(x); \quad \theta > 0.$$

Cari penganggar saksama bervarians minimum secara seragam (PSVMS) untuk median taburan.

[30 markah]

- (d) Andaikan X_1, X_2, \dots, X_n menandai sampel rawak daripada taburan $P_0(\lambda)$ dengan fungsi ketumpatan

$$f(x; \lambda) = \frac{e^{-\lambda} \lambda^x}{x!}, \quad x = 0, 1, 2, \dots$$

Cari batas bawah Cramer-Rao bagi varians penganggar saksama $e^{-\lambda}$.

[30 markah]

4. (a) Assume that \bar{X} denotes the sample mean for a random sample of size n from a $N(\mu, 9)$ distribution. Find the value of n so that $(\bar{X} - 1, \bar{X} + 1)$ is a 95% confidence interval for μ .

[30 marks]

- (b) Assume that X_1, X_2, \dots, X_{10} is a random sample of size 10 from a $N(0, \theta)$ distribution, where $\theta > 0$. For testing $H_0 : \theta = 1$ vs. $H_1 : \theta > 1$, the following critical region is used:

$$C = \left\{ (x_1, x_2, \dots, x_{10}) : \sum_{i=1}^{10} x_i^2 \geq c \right\}.$$

Find c if the size of the critical region C is 0.05.

[30 marks]

- (c) Assume that X is a single observation from a distribution having probability density function

$$f(x; \theta) = (1 + \theta)x^\theta I_{(0,1)}(x); \quad \theta > -1.$$

- (i) Find the most powerful test of size- α for testing $H_0 : \theta = 0$ vs. $H_1 : \theta = 1$.
- (ii) For testing $H_0 : \theta \leq 0$ vs. $H_1 : \theta > 0$, the following test is used:

Reject H_0 if and only if $X \geq \frac{3}{4}$.

Find the power function and the size of the test.

[40 marks]

4. (a) Andaikan bahawa \bar{X} menandakan min sampel bagi suatu sampel rawak bersaiz n daripada taburan $N(\mu, 9)$. Cari nilai n supaya $(\bar{X} - 1, \bar{X} + 1)$ adalah suatu selang keyakinan 95% bagi μ .

[30 markah]

- (b) Andaikan bahawa X_1, X_2, \dots, X_{10} adalah suatu sampel rawak bersaiz 10 daripada taburan $N(0, \theta)$, yang mana $\theta > 0$. Bagi menguji $H_0 : \theta = 1$ lawan $H_1 : \theta > 1$, rantau genting berikut digunakan:

$$C = \left\{ (x_1, x_2, \dots, x_{10}) : \sum_{i=1}^{10} x_i^2 \geq c \right\}.$$

Cari c jika saiz rantau genting C adalah 0.05.

[30 markah]

- (c) Andaikan bahawa X adalah suatu cerapan tunggal daripada taburan yang mempunyai fungsi ketumpatan kebarangkalian

$$f(x; \theta) = (1 + \theta)x^\theta I_{(0,1)}(x); \quad \theta > -1.$$

- (i) Cari ujian paling berkuasa saiz α bagi menguji

$$H_0 : \theta = 0 \text{ vs. } H_1 : \theta = 1.$$

- (ii) Bagi menguji $H_0 : \theta \leq 0$ lawan $H_1 : \theta > 0$, ujian berikut digunakan : Tolak H_0 jika dan hanya jika $X \geq \frac{3}{4}$.

Cari fungsi kuasa dan saiz ujian ini.

[40 markah]

APPENDIX / LAMPIRAN

| Taburan | Fungsi Kumpulan | Min | Varians | Fungsi Penjajana Momen |
|-----------------|---|-------------------------------|--|--|
| Seragam Diskrit | $f(x) = \frac{1}{N} I_{(1,2,\dots,N)}(x)$ | $\frac{N+1}{2}$ | $\frac{N^2-1}{12}$ | $\sum_{j=1}^N \frac{1}{N} e^{j\mu}$ |
| Bernoulli | $f(x) = p^x q^{1-x} I_{(0,1)}(x)$ | p | pq | $q + pe^p$ |
| Binomial | $f(x) = \binom{n}{x} p^x q^{n-x} I_{(0,1,\dots,n)}(x)$ | np | npq | $(q + pe^p)^n$ |
| Geometri | $f(x) = p q^x I_{(0,1,\dots)}(x)$ | $\frac{q}{p}$ | $\frac{q}{p^2}$ | $\frac{p}{1-qe^p}, qe^p < 1$ |
| Poisson | $f(x) = e^{-\lambda} \frac{\lambda^x}{x!} I_{(0,1,\dots)}(x)$ | λ | λ | $\exp\{\lambda(e^p - 1)\}$ |
| Seragam | $f(x) = \frac{1}{b-a} I_{(a,b)}(x)$ | $\frac{a+b}{2}$ | $\frac{(b-a)^2}{12}$ | $\frac{e^{bt} - e^{at}}{(b-a)t}, t \neq 0$ |
| Normal | $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\{-(x-\mu)^2 / 2\sigma^2\} I_{(-\infty,\infty)}(x)$ | μ | σ^2 | $\exp\{\mu t + (\sigma t)^2 / 2\}$ |
| Eksponen | $f(x) = \lambda e^{-\lambda x} I_{(0,\infty)}(x)$ | $\frac{1}{\lambda}$ | $\frac{1}{\lambda^2}$ | $\frac{\lambda}{\lambda-t}, t < \lambda$ |
| Gama | $f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda x} x^{\alpha-1} I_{(0,\infty)}(x)$ | $\frac{\alpha}{\lambda}$ | $\frac{\alpha}{\lambda^2}$ | $\left(\frac{\lambda}{\lambda-t}\right)^\alpha, t < \lambda$ |
| Khi Kuasa Dua | $f(x) = \left(\frac{1}{2}\right)^{r/2} \frac{1}{\Gamma(r/2)} e^{-x/2} x^{(r/2)-1} I_{(0,\infty)}(x)$ | r | $2r$ | $\left(\frac{1}{1-2t}\right)^{r/2}, t < \frac{1}{2}$ |
| Beta | $f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} I_{(0,1)}(x)$ | $\frac{\alpha}{\alpha+\beta}$ | $\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$ | |
| | | | | |